# Mutual Information between Categorical and Gaussian data 

David Atienza

June 7, 2021

## 1 Introduction

This documents shows how to calculate the (conditional) mutual information between categorical and Gaussian data. The mutual information is always calculated between two variables $X$ and $Y$, and can be conditioned on a set of variables $\mathbf{Z}$. Any of this variables can be discrete (categorical) or continuous (Gaussian). When needed, the subscript $D$ and $C$ are used to specify that the variable is discrete or continuous, e.g. $X_{D}$ is a discrete $X$ variable, $\mathbf{Z}_{C}$ is a set of conditioning continuous variables. The instantations for a given variable is denoted using lowercase, e.g. $x_{d}, y_{c}$, etc. The number of categories of the variables $X$ and $Y$ are denoted llx and lly, respectively. The total number of categories for all the $\mathbf{Z}_{D}$ variables is denoted $1 \mathrm{lz}=\prod_{i} l l z_{i}$.

## 2 Mutual Information

### 2.1 Mutual Information between Two Discrete Variables

First we will calculate the mutual information between two discrete variables:

$$
\begin{align*}
I\left(X_{D} ; Y_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)= & H\left(X_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)+H\left(Y_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)-H\left(X_{D}, Y_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)-H\left(\mathbf{Z}_{D}, \mathbf{Z}_{C}\right) \\
= & H\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)+H\left(X_{D}, \mathbf{Z}_{D}\right)+H\left(\mathbf{Z}_{C} \mid Y_{D}, \mathbf{Z}_{D}\right)+H\left(Y_{D}, \mathbf{Z}_{D}\right)  \tag{1}\\
& -H\left(\mathbf{Z}_{C} \mid X_{D}, Y_{D}, \mathbf{Z}_{D}\right)-H\left(X_{D}, Y_{D}, \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{D}\right)
\end{align*}
$$

Note that:

$$
\begin{equation*}
I\left(X_{D} ; Y_{D} \mid \mathbf{Z}_{D}\right)=H\left(X_{D}, \mathbf{Z}_{D}\right)+H\left(Y_{D}, \mathbf{Z}_{D}\right)-H\left(X_{D}, Y_{D}, \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{D}\right) \tag{2}
\end{equation*}
$$

Thus:

$$
\begin{align*}
I\left(X_{D} ; Y_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)= & I\left(X_{D} ; Y_{D} \mid \mathbf{Z}_{D}\right) \\
& +H\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)+H\left(\mathbf{Z}_{C} \mid Y_{D}, \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{C} \mid X_{D}, Y_{D}, \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right) \tag{3}
\end{align*}
$$

where $H\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)$ is the entropy of the Gaussian variables conditioned on $X_{D}, \mathbf{Z}_{D}$. This can be easily calculated:

$$
\begin{align*}
H\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right) & =-\sum_{x_{d} \in X_{D}, \mathbf{z}_{d} \in \mathbf{Z}_{D}} \int f\left(\mathbf{Z}_{C}, x_{d}, \mathbf{z}_{d}\right) \log \left(\frac{f\left(\mathbf{Z}_{C}, x_{d}, \mathbf{z}_{d}\right)}{f\left(x_{d}, \mathbf{z}_{d}\right)}\right) d \mathbf{Z}_{C} \\
& =-\sum_{x_{d} \in X_{D}, \mathbf{z}_{d} \in \mathbf{Z}_{D}} p\left(x_{d}, \mathbf{z}_{d}\right) \int f\left(\mathbf{Z}_{C} \mid x_{d}, \mathbf{z}_{d}\right) \log \left(f\left(\mathbf{Z}_{C} \mid x_{d}, \mathbf{z}_{d}\right)\right) d \mathbf{Z}_{C}  \tag{4}\\
& =\sum_{x_{d} \in X_{D}, \mathbf{z}_{d} \in \mathbf{Z}_{D}} p\left(x_{d}, \mathbf{z}_{d}\right) H\left(\mathbf{Z}_{C} \mid x_{d}, \mathbf{z}_{d}\right)
\end{align*}
$$

The last integral in (4) is the entropy of the Gaussian distribution trained with the $x_{d}, \mathbf{z}_{d}$ configuration.

For a Gaussian distribution, this integral can be solved with a closed-form formula:

$$
\begin{equation*}
H\left(\mathbf{Z}_{C} \mid x_{d}, \mathbf{z}_{d}\right)=\frac{k}{2}+\frac{k}{2} \log (2 \pi)+\frac{1}{2} \log \left(\left|\boldsymbol{\Sigma}_{x_{d}, \mathbf{z}_{d}}\right|\right) \tag{5}
\end{equation*}
$$

where $k=\left|\mathbf{Z}_{C}\right|$ is the dimensionality of the Gaussian distribution and $\boldsymbol{\Sigma}_{x_{d}, \mathbf{Z}_{d}}$ is the covariance of the data with the discrete configuration $x_{d}, \mathbf{z}_{d}$.

The remaining terms $H\left(\mathbf{Z}_{C} \mid Y_{D}, \mathbf{Z}_{D}\right), H\left(\mathbf{Z}_{C} \mid X_{D}, Y_{D}, \mathbf{Z}_{D}\right)$ and $H\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)$ can be calculated similarly.

### 2.2 Mutual Information between a Discrete and Continuous Variable

$$
\begin{align*}
I\left(X_{D} ; Y_{C} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)= & H\left(X_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)+H\left(Y_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)-H\left(X_{D}, Y_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)-H\left(\mathbf{Z}_{D}, \mathbf{Z}_{C}\right) \\
= & H\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)+H\left(X_{D}, \mathbf{Z}_{D}\right)+H\left(Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)+H\left(\mathbf{Z}_{D}\right) \\
& -H\left(Y_{C}, \mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)-H\left(X_{D}, \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{D}\right) \\
= & H\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)+H\left(Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)-H\left(Y_{C}, \mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right) \tag{6}
\end{align*}
$$

where the entropy terms can be calculated as in (4), but in this case the variable $Y$ is added in some of the estimated multivariate Gaussian distributions.

### 2.3 Mutual Information between Two Continuous Variable

For an unconditional mutual information, the mutual information can be calculated with the correlation coefficient:

$$
\begin{equation*}
I\left(X_{C} ; Y_{C}\right)=-\frac{1}{2} \log \left(1-\rho^{2}\right) \tag{7}
\end{equation*}
$$

where $\rho$ is the linear correlation coefficient between $X$ and $Y$.
For the general case:

$$
\begin{align*}
I\left(X_{C} ; Y_{C} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)= & H\left(X_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)+H\left(Y_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)-H\left(X_{C}, Y_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)-H\left(\mathbf{Z}_{D}, \mathbf{Z}_{C}\right) \\
= & H\left(X_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)+H\left(\mathbf{Z}_{D}\right)+H\left(Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)+H\left(\mathbf{Z}_{D}\right) \\
& -H\left(X_{C}, Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{D}\right) \\
= & H\left(X_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)+H\left(Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)-H\left(X_{C}, Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)-H\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right) \tag{8}
\end{align*}
$$

where the entropy terms can be calculated as in (4), but in this case the variables $X$ and $Y$ are added in some of the estimated multivariate Gaussian distributions.

## 3 Empirical Degrees of Freedom

This section shows the empirical degrees of freedom by running a simulation over 1000 datasets of 100000 instances that are compatible with the null hypothesis (conditional independence). The empirical degrees of freedom have been rounded to the nearest integer number.

### 3.1 Empirical Degrees of Freedom between Two Discrete Variables

|  | nt. | par |  | 2 cont. parents |  |  |  | 3 cont. parents |  |  |  | 4 cont. parents |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 11 x | lly | llz | df | llx | lly | 1 lz | df | 1lx | lly | 1 lz | df | llx | lly | 11 z | df |
| 2 | 4 | 3 | 18 | 2 | 4 | 3 | 36 | 2 | 4 | 3 | 63 | 2 | 4 | 3 | 99 |
| 2 | 3 | 4 | 16 | 2 | 3 | 4 | 32 | 2 | 3 | 4 | 56 | 2 | 3 | 4 | 88 |
| 3 | 4 | 2 | 24 | 3 | 4 | 2 | 48 | 3 | 4 | 2 | 84 | 3 | 4 |  | 132 |

Inducted formula:

$$
\begin{equation*}
\mathrm{df}=(\mathrm{llx}-1) \cdot(\mathrm{lly}-1) \cdot \mathrm{llz} \cdot\left[1+\frac{\left|\mathbf{Z}_{C}\right| \cdot\left(\left|\mathbf{Z}_{C}\right|+1\right)}{2}\right] \tag{9}
\end{equation*}
$$

### 3.2 Empirical Degrees of Freedom between a Discrete and Continuous Variable

| 1 cont. parents |  |  | 2 cont. parents |  |  | 3 cont. parents |  |  | 4 cont. parents |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| llx | llz | df | llx | 1 lz | df | llx | llz | df | llx | 1 lz | df |
| 2 | 3 | 6 | 2 | 3 | 9 | 2 | 3 | 12 | 2 | 3 | 15 |
| 2 | 4 | 8 | 2 | 4 | 12 | 2 | 4 | 16 | 2 | 4 | 20 |
| 3 | 2 | 8 | 3 | 2 | 12 | 3 | 2 | 16 | 3 | 2 | 20 |
| 3 | 4 | 16 | 3 | 4 | 24 | 3 | 4 | 32 | 3 | 4 | 40 |
| 4 | 2 | 12 | 4 | 2 | 18 | 4 | 2 | 24 | 4 | 2 | 30 |
| 4 | 3 | 18 | 4 | 3 | 27 | 4 | 3 | 36 | 4 | 3 | 45 |

Inducted formula:

$$
\begin{equation*}
\mathrm{df}=(\mathrm{llx}-1) \cdot \mathrm{llz} \cdot\left[1+\left|\mathbf{Z}_{C}\right|\right] \tag{10}
\end{equation*}
$$

### 3.3 Empirical Degrees of Freedom between Two Continuous Variables

 Inducted formula:$$
\begin{equation*}
\mathrm{df}=\mathrm{llz} \tag{11}
\end{equation*}
$$

1 cont. parents
contx conty llz df

| 2 | 2 |
| :--- | :--- |
| 3 | 3 |
| 4 | 4 |

2 cont. parents contx conty llz df $2 \quad 2$
33
$4 \quad 4$

(a) $H_{0}$ model

3 cont. parents
contx conty llz df $2 \quad 2$
33
44

(b) $H_{1}$ model

Figure 1: Null hypothesis (left) and alternative (right) models for two discrete variables.

## 4 Asymptotic Degrees of Freedom

### 4.1 Asymptotic Degrees of Freedom between Two Discrete Variables

There is a direct relationship between mutual information and a likelihood ratio test ( $G$-test):

$$
\begin{equation*}
G=2 \cdot N \cdot I\left(X_{D} ; Y_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right) \tag{12}
\end{equation*}
$$

The $G$ statistic is distributed as a $\chi^{2}$ if the null hypothesis is true. The degrees of freedom of the $\chi^{2}$ distribution is the difference in the number of free parameters between a model where there is conditional dependence between $X_{D}$ and $Y_{D}$ (Figure 1b), and a model where there is no conditional dependence between $X_{D}$ and $Y_{D}$ (Figure 1a).

The only node that contains a different number of parameters is the conditional distribution of $Y_{D}$. So we must analyze that distribution to find the degrees of freedom of the $\chi^{2}$ distribution.

For the $H_{0}$ model, the distribution $f\left(Y_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ can be defined using the Bayes rule (as in a conditional linear Gaussian networks the discrete nodes do not have continuous parents):

$$
\begin{equation*}
f\left(Y_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)=\frac{f\left(\mathbf{Z}_{C} \mid Y_{D}, \mathbf{Z}_{D}\right) f\left(Y_{D} \mid \mathbf{Z}_{D}\right)}{f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)} \tag{13}
\end{equation*}
$$

Note that only lly -1 models are needed to be fitted because the probabilities $f\left(Y_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ must sum to 1 , so the probability for the last category can be defined as:

$$
\begin{equation*}
f\left(Y_{D}=\text { lly } \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)=1-\sum_{i=1}^{\mathrm{lly}-1} \frac{f\left(\mathbf{Z}_{C} \mid Y_{D}=i, \mathbf{Z}_{D}\right) f\left(Y_{D}=i \mid \mathbf{Z}_{D}\right)}{f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)} \tag{14}
\end{equation*}
$$

$f\left(\mathbf{Z}_{C} \mid Y_{D}, \mathbf{Z}_{D}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
(\mathrm{lly}-1) \cdot \mathrm{llz} \cdot \frac{\left|\mathbf{Z}_{C}\right| \cdot\left(\left|\mathbf{Z}_{C}\right|+3\right)}{2} \tag{15}
\end{equation*}
$$

Note that $\frac{\left|\mathbf{Z}_{C}\right| \cdot\left(\left|\mathbf{Z}_{C}\right|+3\right)}{2}$ is the number of free parameters of a multivariate Gaussian distribution of $\left|\mathbf{Z}_{C}\right|$ dimensions.
$f\left(Y_{D} \mid \mathbf{Z}_{D}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
(\mathrm{lly}-1) \cdot \mathrm{llz} \tag{16}
\end{equation*}
$$

$f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)$ does not contain free parameters because it can be represented using the previous functions:

$$
\begin{equation*}
f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)=\sum_{y_{d} \in Y_{D}} f\left(\mathbf{Z}_{C} \mid Y_{D}=y_{d}, \mathbf{Z}_{D}\right) f\left(Y_{D}=y_{d} \mid \mathbf{Z}_{D}\right) \tag{17}
\end{equation*}
$$

For the $H_{1}$ model, the distribution $f\left(Y_{D} \mid X_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ can be defined using the Bayes rule (as in a conditional linear Gaussian networks the discrete nodes do not have continuous parents):

$$
\begin{equation*}
f\left(Y_{D} \mid X_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)=\frac{f\left(\mathbf{Z}_{C} \mid X_{D}, Y_{D}, \mathbf{Z}_{D}\right) f\left(Y_{D} \mid X_{D}, \mathbf{Z}_{D}\right)}{f\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)} \tag{18}
\end{equation*}
$$

Note that only lly-1 models are needed to be fitted because the probabilities $f\left(Y_{D} \mid X_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ must sum to 1 , so the probability for the last category can be defined as:

$$
\begin{equation*}
f\left(Y_{D}=\text { lly } \mid X_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)=1-\sum_{i=1}^{\text {lly }-1} \frac{f\left(\mathbf{Z}_{C} \mid X_{D}, Y_{D}=i, \mathbf{Z}_{D}\right) f\left(Y_{D}=i \mid X_{D}, \mathbf{Z}_{D}\right)}{f\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)} \tag{19}
\end{equation*}
$$

$f\left(\mathbf{Z}_{C} \mid X_{D}, Y_{D}, \mathbf{Z}_{D}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
(\mathrm{lly}-1) \cdot \mathrm{llx} \cdot \mathrm{llz} \cdot \frac{\left|\mathbf{Z}_{C}\right| \cdot\left(\left|\mathbf{Z}_{C}\right|+3\right)}{2} \tag{20}
\end{equation*}
$$

$f\left(Y_{D} \mid X_{D}, \mathbf{Z}_{D}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
(l l y-1) \cdot l l x \cdot l l z \tag{21}
\end{equation*}
$$

$f\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)$ does not contain free parameters because it can be represented using the previous functions:

$$
\begin{equation*}
f\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)=\sum_{y_{d} \in Y_{D}} f\left(\mathbf{Z}_{C} \mid X_{D}, Y_{D}=y_{d}, \mathbf{Z}_{D}\right) f\left(Y_{D}=y_{d} \mid X_{D}, \mathbf{Z}_{D}\right) \tag{22}
\end{equation*}
$$

The difference in parameters (and the degrees of freedom of the $\chi^{2}$ ) is equal to:

$$
\begin{align*}
\mathrm{df} & =\mathrm{lly}-1) \cdot(\mathrm{llx}-1) \cdot \mathrm{llz} \cdot \frac{\left|\mathbf{Z}_{C}\right| \cdot\left(\left|\mathbf{Z}_{C}\right|+3\right)}{2}+(\mathrm{lly}-1) \cdot(\mathrm{llx}-1) \cdot \mathrm{llz} \\
& =(\mathrm{lly}-1) \cdot(\mathrm{llx}-1) \cdot \mathrm{llz}\left[1+\frac{\left|\mathbf{Z}_{C}\right| \cdot\left(\left|\mathbf{Z}_{C}\right|+3\right)}{2}\right] \tag{23}
\end{align*}
$$

### 4.2 Asymptotic Degrees of Freedom between a Discrete and Continuous Variable

The $H_{0}$ model is shown in Figure 2a and the $H_{1}$ model is shown in Figure 2b.
The distribution $f\left(Y_{C} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ has a number of free parameters equal to:

(a) $H_{0}$ model

(b) $H_{1}$ model

Figure 2: Null hypothesis (left) and alternative (right) models for a continuous and discrete variable.

(a) $H_{0}$ model

(b) $H_{1}$ model

Figure 3: Variation of the null hypothesis (left) and alternative (right) models for a continuous and discrete variable.

$$
\begin{equation*}
\mathrm{llz} \cdot\left(\left|\mathbf{Z}_{C}\right|+2\right) \tag{24}
\end{equation*}
$$

The distribution $f\left(Y_{C} \mid X_{D}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
\mathrm{llx} \cdot \mathrm{llz} \cdot\left(\left|\mathbf{Z}_{C}\right|+2\right) \tag{25}
\end{equation*}
$$

The difference in parameters (and the degrees of freedom of the $\chi^{2}$ ) is equal to:

$$
\begin{equation*}
\mathrm{df}=(\mathrm{llx}-1) \cdot \mathrm{llz} \cdot\left(\left|\mathbf{Z}_{C}\right|+2\right) \tag{26}
\end{equation*}
$$

The same result can be derived using a different $H_{0}$ model (Figure 3a) and $H_{1}$ model (Figure 3b). In this case, the difference in the number of parameters happens to be in the conditional distribution of $X_{D}$.

The $f\left(X_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ can be defined using the Bayes rule:

$$
\begin{equation*}
f\left(X_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)=\frac{f\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right) f\left(X_{D} \mid \mathbf{Z}_{D}\right)}{f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)} \tag{27}
\end{equation*}
$$

Note that only llx -1 models are needed to be fitted because the probabilities $f\left(X_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ must sum to 1 , so the probability for the last category can be defined as:

$$
\begin{equation*}
f\left(X_{D}=\operatorname{llx} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)=1-\sum_{i=1}^{\mathrm{llx}-1} \frac{f\left(\mathbf{Z}_{C} \mid X_{D}=i, \mathbf{Z}_{D}\right) f\left(X_{D}=i \mid \mathbf{Z}_{D}\right)}{f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)} \tag{28}
\end{equation*}
$$

The distribution $f\left(\mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
(\mathrm{llx}-1) \cdot \mathrm{llz} \cdot \frac{\left|\mathbf{Z}_{C}\right| \cdot\left(\left|\mathbf{Z}_{C}\right|+3\right)}{2} \tag{29}
\end{equation*}
$$

The distribution $f\left(X_{D} \mid \mathbf{Z}_{D}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
(1 l x-1) \cdot l l z \tag{30}
\end{equation*}
$$

$f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)$ does not contain free parameters because it can be represented using the previous functions:

$$
\begin{equation*}
f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)=\sum_{x_{d} \in X_{D}} f\left(\mathbf{Z}_{C} \mid X_{D}=x_{d}, \mathbf{Z}_{D}\right) f\left(X_{D}=x_{d} \mid \mathbf{Z}_{D}\right) \tag{31}
\end{equation*}
$$

The $f\left(X_{D} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ can be defined using the Bayes rule:

$$
\begin{equation*}
f\left(X_{D} \mid Y_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)=\frac{f\left(Y_{C}, \mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right) f\left(X_{D} \mid \mathbf{Z}_{D}\right)}{f\left(Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)} \tag{32}
\end{equation*}
$$

Note that only llx - 1 models are needed to be fitted because the probabilities $f\left(X_{D} \mid\right.$ $\left.Y_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ must sum to 1 , so the probability for the last category can be defined as:

$$
\begin{equation*}
f\left(X_{D}=\mathrm{llx} \mid Y_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)=1-\sum_{i=1}^{\mathrm{lnx}-1} \frac{f\left(Y_{C}, \mathbf{Z}_{C} \mid X_{D}=i, \mathbf{Z}_{D}\right) f\left(X_{D}=i \mid \mathbf{Z}_{D}\right)}{f\left(Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)} \tag{33}
\end{equation*}
$$

The distribution $f\left(Y_{C}, \mathbf{Z}_{C} \mid X_{D}, \mathbf{Z}_{D}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
(\mathrm{llx}-1) \cdot \mathrm{llz} \cdot \frac{\left(\left|\mathbf{Z}_{C}\right|+1\right) \cdot\left(\left|\mathbf{Z}_{C}\right|+4\right)}{2} \tag{34}
\end{equation*}
$$

The distribution $f\left(X_{D} \mid \mathbf{Z}_{D}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
(1 l x-1) \cdot l l z \tag{35}
\end{equation*}
$$

$f\left(\mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)$ does not contain free parameters because it can be represented using the previous functions:

$$
\begin{equation*}
f\left(Y_{C}, \mathbf{Z}_{C} \mid \mathbf{Z}_{D}\right)=\sum_{x_{d} \in X_{D}} f\left(Y_{C}, \mathbf{Z}_{C} \mid X_{D}=x_{d}, \mathbf{Z}_{D}\right) f\left(X_{D}=x_{d} \mid \mathbf{Z}_{D}\right) \tag{36}
\end{equation*}
$$

The difference in parameters (and the degrees of freedom of the $\chi^{2}$ ) is equal to:

$$
\begin{align*}
\mathrm{df} & =(\mathrm{llx}-1) \cdot \mathrm{llz} \cdot\left(\frac{\left|\mathbf{Z}_{C}\right|^{2}+5\left|\mathbf{Z}_{C}\right|+4-\left|\mathbf{Z}_{C}\right|^{2}-3\left|\mathbf{Z}_{C}\right|}{2}\right) \\
& =(\mathrm{llx}-1) \cdot \mathrm{llz} \cdot\left(\frac{2\left|\mathbf{Z}_{C}\right|+4}{2}\right)  \tag{37}\\
& =(1 \mathrm{~lx}-1) \cdot \mathrm{llz} \cdot\left(\left|\mathbf{Z}_{C}\right|+2\right)
\end{align*}
$$


(a) $H_{0}$ model

(b) $H_{1}$ model

Figure 4: Null hypothesis (left) and alternative (right) models for two continuous variables.

### 4.3 Asymptotic Degrees of Freedom between Two Continuous Variables

The $H_{0}$ model is shown in Figure 4a and the $H_{1}$ model is shown in Figure 4b.
For the $H_{0}$ model, the distribution $f\left(Y_{C} \mid \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
\mathrm{llz} \cdot\left(\left|\mathbf{Z}_{C}\right|+2\right) \tag{38}
\end{equation*}
$$

For the $H_{1}$ model, the distribution $f\left(Y_{C} \mid X_{C}, \mathbf{Z}_{D}, \mathbf{Z}_{C}\right)$ has a number of free parameters equal to:

$$
\begin{equation*}
\mathrm{llz} \cdot\left(\left|\mathbf{Z}_{C}\right|+3\right) \tag{39}
\end{equation*}
$$

The difference in parameters (and the degrees of freedom of the $\chi^{2}$ ) is equal to:

$$
\begin{array}{|l|l|}
\hline 1 \mathrm{l}  \tag{40}\\
\hline
\end{array}
$$

